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ON A CHARACTERIZATION OF THE EXPONENTIAL
DISTRIBUTION BY ORDER STATISTICS

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O Instituto é o organismo especializado da OEA para o setor agropecuário. Foi estabelecido em 1942 pelos governos americanos com o propósito de ajudar os países a estimularem e promoverem o desenvolvimento rural, como meio para alcançar o desenvolvimento geral e o bem-estar da população.

0. SUMMARY

A characterization of the exponential distribution is shown by considering the identical distribution of the random variables $n X_{1,n}$ and $(n-i+1) (X_{i,n} - X_{i-1,n})$, where $X_{j,n}$ is the j th order statistic, $(j=1, i, i-1)$ is a sample of size n .

1. INTRODUCTION

Let be a random variable (r.v) whose probability density function (p.d.f.) is

$$f(x) = \theta^{-1} \exp(-x/\theta), \quad x > 0 \quad \theta > 0$$

$$= 0, \text{ otherwise.} \quad (1.1)$$

Suppose X_1, X_2, \dots, X_n be a random sample of size n from a population with p.d.f. $f(x)$, $x > 0$, and let $X_{1,n} < X_{2,n} < \dots < X_{n,n}$ be the associated order statistics. Kotz (1974) and Galambos (1975) discussed extensively the characterization of exponential distribution by order statistics. Desu (1971) showed that the exponential distribution is the only non-degenerate one that for all K , K times the minimum of the random sample of size K from the distribution has the same distribution as a single observation. Sethuraman (1965) proved a stronger form of Desu's theorem. Puri and Rubin (1970) proved that if X_1, X_2 are independent copies of a r.v. X with p.d.f. $f(x)$, then X and $|X_1 - X_2|$ have the same distribution if and only if $f(x)$ as given in (1.1). If $F(x)$ has the density $f(x)$, the ratio $h(x) = f(x) / (1 - F(x))$, for $F(x) < 1$, is called the hazard rate. Most materials, structures and

devices where replacement is considered wear out with time; the distribution with increasing hazard rate are of interest.

The distribution with decreasing hazard rate may also arise in reliability as reflection of 'work hardening' etc. Barlow, Marshall and Proschan (1963) considered properties of probability distribution with monotone hazard rate. Ahsanullah (1975) gave a characterization of exponential distribution by considering identical distribution of $n X_{1,n}$ and $X_{n,n} - X_{n-1,n}$. In this paper a generalisation of the author's (1975) work is given.

2. CHARACTERIZATION

Let X be a non-negative r.v. having an absolutely continuous (with respect to lebesgue measure) strictly increasing distribution function $F(x)$ for all $x > 0$, and $F(x) < 1$, for all x . Then the following properties are equivalent:

- (a) X has an exponential distribution with density as given in (1.1)
- (b) X has a monotone hazard rate and for any fixed i , and n , $2 \leq i \leq n$, the statistics $(n-i+1) (X_{i,n} - X_{i-1,n})$ and $n X_{1,n}$ are identically distributed.

Proof: (a) \implies (b)

Let $h(x) = f(x)/(1-F(x)) = \theta^{-1}$. Considering the joint probability of $X_{i,n}$ and $X_{i-1,n}$ and using the transformation $U = X_{i,n}$ and $V = (n-i+1) (X_{i,n} - X_{i-1,n})$ it can be shown, using the value of $h(x) = \theta^{-1}$, that V is identically distributed as $n X_{1,n}$.

Proof: (b) \implies (a)

Let $Y_i = X_{i,n}$, the probability density of Y_i , Pyke (1965), is

$$f_{Y_i}(y) = \frac{n!}{(i-2)!(n-i)!} \int_0^\infty (F(u))^{i-2} (1-F(u+y))^{n-i} f(u) f(u+y) du \quad (2.1)$$

Substituting $Z = (n-i+1) Y_i$, we get the p.d.f. of Z , as

$$f_Z(z) = \frac{n!}{(i-2)!(n-i+1)!} \int_0^\infty (F(u))^{i-2} (1-F(u+z/(n-i+1)))^{n-i} f(u) f(u+z/(n-i+1)) du \quad (2.2)$$

The probability density function of $W = nX_{i,n}$ is

$$f_W(w) = (1-F(w/n))^{n-1} f(w/n)$$

Since Z and W are identically distributed, using the fact

$$\frac{(i-2)!(n-i+1)!}{n!} = \int_0^\infty (F(u))^{i-2} (1-F(u))^{n-i+1} f(u) du,$$

We get on simplification from (2.2) and (2.3),

$$0 = \int_0^\infty (F(u))^{i-2} (1-F(u))^{n-i+1} f(u) g(u,z) du, \text{ for all } z, \quad (2.4)$$

Where

$$g(u, z_1) = \frac{(1-F(z/n))^{n-1} f(z/n) - (1-F(u+z/(n-i+1)))^{n-i} f(u+z/(n-i+1))}{(1-F(u))^{n-i+1}}$$

Integrating (2.4), we respect to z from 0 to z_1 , we obtain after some simplification

$$0 = \int_0^\infty (F(u))^{i-2} (1-F(u))^{n-i+1} f(u) G(u, z_1) du = 0, \text{ for all } z_1 \quad (2.5)$$

where

$$G(u, z_1) = \left(\frac{1-F(u + \frac{z_1}{n-i+1})}{1-F(u)} \right)^{n-i+1} - (1-F(\frac{z_1}{n}))^n \quad (2.6)$$

By lemma 1, given in appendix, we have

$$G(0, z_1) = 0, \text{ for all } z_1, i, e,$$

$$(1-F(\frac{z}{n-i+1}))^{n-i+1} = (1-F(\frac{z}{n}))^n, \text{ for all } z \quad (2.7)$$

Substituting $H(z) = 1-F(z)$, $\Psi(z) = -\log H(z)$ and

$$y = \frac{1}{n} z, \text{ we get}$$

$$\Psi(y) = \frac{n-i+1}{n} \Psi(\frac{n}{n-i+1} y), \text{ for all } y, \quad (2.8)$$

We know $F(z)$ is strictly monotone increasing function with $F(z) < 1$, for all z , hence $y(z)$ is continuous for all z , thus the solution of (2.8) Aczel (1966), is

$$\Psi(y) = cy \quad (2.9)$$

Where c is constant, thus $F(y) = 1-e^{cy}$

Using the boundary condition $F(0) = 0$, and $F(\infty) = 1$,

we get

$$F(x) = 1-e^{-\theta x} \quad (2.10)$$

Where $\theta > 0$

3. APPENDIX

Lemma I:

Let $F(x)$ be the absolutely continuous (with respect to lebesgue measure) monotone strictly increasing distribution function on a non-negative r.v. X , and $F(0) = 0$, $F(x) < 1$, for all x and $f(x)$ be the corresponding p.d.f.

If the hazard rate is monotone and if

$$\int_0^{\infty} (F(u))^{i-2} (1-F(u))^{n-i+1} f(u) G(u, z) du = 0, \quad (A.1)$$

for fixed i , and n and all z , where

$$G(u, z) = \left(\frac{1-F(u + \frac{z}{n-i+1})}{1-F(u)} \right)^{n-i+1} - (1-F(\frac{z}{n}))^n,$$

$G(0, z) = 0$, for all z .

Proof:

We have $G(u, 0) = 0 = G(0, \infty)$

$$\frac{\partial}{\partial u} G(u, z) = (n-i+1) q(u, z) (h(u) - h(u + \frac{z}{n-i+1})) \quad (A.2)$$

where

$$q(u, z) = \left(\frac{1-F(u + \frac{z}{n-i+1})}{1-F(u)} \right)^{n-i+1}$$

(i) $h(x) = \text{constant}$ for all x .

$\frac{\partial}{\partial u} G(u, z) = 0$, for all u and z , hence $G(u, z)$ is independent of u for all u and z .

(ii) $h(x)$ is strictly monotone decreasing in x .

It follows from (A.2) that $G(u, z)$ is strictly increasing in u , with fixed z . For (A.1) to be true $G(u, z_1)$ for $z=z_1$, if $G(u, z_1) \neq 0$, must be negative in an interval including the point zero. Let $u \in I$, where I is an interval $[0, b]$, b is a real number. Then $G(u_1, z_1) < 0$, for $u \in I$.

$$\frac{\partial}{\partial z_1} G(u_1, z_1) = (1-F(z_1/n))^n \ell(u_1, z_1)$$

where

$$\ell(u_1, z_1) = h(z_1) - m(u_1, z_1) h(u_1 + z_1/(n-i+1))$$

and

$$m(u_1, z_1) = (1/(1-F(z_1/n)))^n ((1-F(u_1 + z_1/(n-i+1)))/(1-F(u_1)))^n$$

$$G(u_1, z_1) < 0 \Rightarrow m(u_1, z_1) < 1$$

Hence for $u \in I$,

$$\begin{aligned} \frac{\partial}{\partial z_1} G(u_1, z_1) &> (1-F(z_1/n))^n (h(z_1) - h(u_1 + z_1/(n-i+1))) \\ &> 0 \end{aligned}$$

Thus $G(u_1, z_1)$ is increasing with z_1 . But $G(u_1, 0)=0$

and $G(u_1, \infty)=0$, therefore it follows by the continuity of $G(u_1, z_1)$

that $G(u_1, z_1)=0$, for all z_1 . Hence $G(0, z_1)=0$, for all z_1 .

(iii) $h(x)$ is strictly monotone increasing in x . In this case $G(u, z)$ will be strictly monotone decreasing in u with fixed z . Following the similar argument as in (ii) it can be shown that $G(0, z)=0$ for all z .

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