

Genetic analysis of productive and reproductive traits in multiple-breed dairy cattle populations

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Abstract – The objective of this work was to determine whether the random regression model using linear splines (RRMLS) is suitable to estimate the genetic parameters for productive and reproductive traits of a multiple-breed dairy cattle population, as well as to investigate the effect of the genetic group of the progeny on the genetic merit of the sire. The multiple-trait model (MTM) and the RRMLS with one knot fitted for every genetic group were used to obtain the genetic parameters. Records of 1/2 Holstein + 1/2 Gyr (1/2HG), 5/8 Holstein + 3/8 Gyr (5/8HG), and 3/4 Holstein + 1/4 Gyr (3/4HG) crossbreed dams were considered. The RRMLS showed better fitting. The additive and residual variances estimated by the MTM and the RRMLS were similar. Heritability varied from 0.20 to 0.33 for age at first calving, from 0.09 to 0.22 for first lactation length, and from 0.15 to 0.35 for first lactation 305-day milk yield, according to the genetic composition of the dams. The RRMLS is suitable to estimate the genetic parameters for productive and reproductive traits of multiple-breed dairy cattle populations. The genetic merit of the sires is affected by the genetic group of the progeny by which they are evaluated.

Index terms: Gyr, heritability, heterosis, Holstein, selection.

Análise genética de características produtivas e reprodutivas em populações multirraciais de bovinos leiteiros

Resumo – O objetivo deste trabalho foi verificar se o modelo de regressão aleatória com polinômios lineares do tipo spline (MRAPLS) é adequado para estimar os parâmetros genéticos de características produtivas e reprodutivas em uma população multirracial de bovinos leiteiros, bem como investigar o efeito do grupo genético da progênie sobre o mérito genético do touro. Foram utilizados o modelo multicaracterístico (Mult) e o MRAPLS com nós ajustados aos grupos genéticos, para obtenção dos parâmetros genéticos. Foram considerados os registros de vacas 1/2 Holandesa + 1/2 Gir (1/2HG), 5/8 Holandesa + 3/8 Gir (5/8HG) e 3/4 Holandesa + 1/4 Gir (3/4HG). O MRAPLS apresentou os melhores ajustes. As variâncias aditivas e residuais estimadas por Mult e MRAPLS foram semelhantes. As herdabilidades variaram de 0,20 a 0,33 para idade ao primeiro parto, de 0,09 a 0,22 para duração da primeira lactação e de 0,15 a 0,35 para produção de leite em até 305 dias na primeira lactação, de acordo com a composição genética das vacas. O MRAPLS é adequado para estimar parâmetros genéticos de características produtivas e reprodutivas em populações multirraciais de bovinos leiteiros. O mérito genético do touro é influenciado pelo grupo genético da progênie por meio do qual ele é avaliado.

Termos para indexação: Gir, herdabilidade, heterose, Holandesa, seleção.

Introduction

The success of a genetic breeding program depends on the definition of the selection goals and on the choice

of the most adequate genotypes for the production system. The practice of crossing different cattle breeds aiming milk yield has increased due to the possibility of improving reproductive (Dhakal et al., 2013; Yao et

al., 2014) and adaptive traits (Berman, 2011). These crossings also contribute to increase the productive and economic performance of the production systems (Wolfová et al., 2007; Dal Zotto et al., 2009).

For more accurate genetic evaluations, differences in breed composition, in terms of additive and non-additive effects and of heterogeneous variances, should be taken into account (Cardoso & Tempelman, 2004; VanRaden et al., 2007). In addition, the interactions between the breeding value of the bull and the genetic group its progenies belong to are important to establish breeding strategies aiming to obtain greater benefits for future generations (Vallée et al., 2013; Lukaszewicz et al., 2015). Therefore, the genetic evaluation of multiple-breed populations may be improved when the traits of each genetic group are specifically assessed.

Multiple-trait models (MTM) allow considering any characteristic (e.g., milk yield) as a distinct one in each group it is measured in. This approach has already been adopted in other studies (Vallée et al., 2013; Lukaszewicz et al., 2015) and allows detecting differences related to variances and covariances. However, these models tend to become more parameterized as the number of genetic groups increases. This parametrization may be reduced by using random regression models for multiple-breed genetic evaluations.

The use of a random regression model using linear splines (RRMLS) is an alternative for cases in which the MTM shows limitations. Misztal (2006) pointed out that polynomial splines can be used to adjust the effect of a continuous variable, such as the expected percentage of alleles of a certain breed. The author also found that the (co)variances of the polynomial random regression coefficients are equivalent to those obtained for the traits evaluated in the levels defined by the knot polynomials. That is, if the knots are positioned along the possible percentage values for the alleles of a certain breed (e.g., 0.0, 0.5, 0.75, and 1.0), the (co)variances of the random regression coefficients will be equivalent to the those between the trait evaluated in each animal and the breed group (in this case, 0.0, 0.5, 0.75, and 1.0). In the RRMLS, the knots can be positioned in points of interest or based on the statistical criteria analysis. Therefore, this model may be less parameterized than the MTM, when the (co) variances for some intermediary groups (with a small

number of observations) may be obtained through interpolation.

The objective of this work was to determine whether the RRMLS is suitable to estimate the genetic parameters for productive and reproductive traits of a multiple-breed dairy cattle population, as well as to investigate the effect of the genetic group of the progeny on the breeding value of the bull.

Materials and Methods

The phenotypic data were obtained from the milk production services of Associação Brasileira dos Criadores de Girolando and are part of the Brazilian archive on dairy cattle of Embrapa Gado de Leite. The assessed traits were: age at first calving (AFC), first lactation length (FLL), and first lactation 305-day milk yield (305MY). Records of 1/2 Holstein + 1/2 Gyr (1/2HG), 5/8 Holstein + 3/8 Gyr (5/8HG), and 3/4 Holstein + 1/4 Gyr (3/4HG) crossbred dams were used (Table 1).

The records considered were those of dams with AFC from 560 to 1,650 days, born from 1997 to 2010, which calved from 1998 to 2013. Data on lactations terminated due to low yield or to the start of calving were also taken into account, being considered within normality. Only data on contemporary groups with, at least, three records and female progenies of at least two sires were considered (Table 2). The contemporary groups were formed by the variables herd and birth year for AFC and herd and calving year for FLL and 305MY.

A MTM with three traits was used to analyze AFC, FLL, and 305MY. This way, the same one traits evaluated in the 1/2HG, 5/8HG, and 3/4HG genetic groups was considered as three distinct ones. The MTM may be represented by:

$$y_{ijklmn} = cg_i + m_j + b_1PHet_k + b_2PHol_l + b_3age_m + b_4age_m^2 + a_n + e_{ijklmn}$$

where y_{ijklmn} is the phenotype of the traits; cg_i is the effect of contemporary group i ; m_j is the effect of month j (calving month for FLL and 305MY or birth month for AFC); b_1 , b_2 , b_3 , and b_4 are the regression coefficients that adjust the heterozygosis effects, the expected percentage of alleles of the Holstein breed and dam age at calving (linear and quadratic, respectively); $PHet_k$ and $PHol_l$ are the expected percentage of

heterozygosity k and of alleles of the Holstein breed l , respectively; age_m represents dam age at calving m ; a_n is the additive genetic value of animal n ; and e_{ijklmn} is the error associated with each observation. The effects of dam age at calving (linear and quadratic) were adjusted only for the traits FLL and 305MY.

The PHol of each animal was obtained from the average PHol of its parents, based on their oldest ancestors. The PHet was determined using the equation: $PHet = PHol_s \times (1 - PHol_d) + PHol_d \times (1 - PHol_s)$ where PHet is the percentage of heterozygosity of the individual of interest; PHol_s is the expected percentage of alleles of the female Holstein parent; and PHol_d is the expected percentage of alleles of the male Holstein parent.

The MTM model may be represented by the following matrix:

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \end{bmatrix} = \begin{bmatrix} X_1 & \Phi & \Phi \\ \Phi & X_2 & \Phi \\ \Phi & \Phi & X_3 \end{bmatrix} \begin{bmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix} + \begin{bmatrix} Z_1 & \Phi & \Phi \\ \Phi & Z_2 & \Phi \\ \Phi & \Phi & Z_3 \end{bmatrix} \begin{bmatrix} a_1 \\ a_2 \\ a_3 \end{bmatrix} + \begin{bmatrix} e_1 \\ e_2 \\ e_3 \end{bmatrix}$$

where y_h are the vector observations of trait h ($h = 1, 2, \text{ and } 3$); X_h is the incidence matrix of the effects of

trait h ; Φ is a null matrix; β_h is the solution vector for fixed effects; Z_h is the incidence matrix of the direct additive genetic effects; a_h is the solution vector for the direct additive genetic effects; and e_h is the error vector.

The trait AFC, FLL, and 305MY were also evaluated through the RRMLS. In this model, knots were positioned at the 0.5, 0.625, and 0.75 points, which correspond to the expected percentage of the alleles of the Holstein breed in the 1/2HG, 5/8HG, and 3/4HG genetic groups. The residual variances were considered heterogeneous, and a variance class was for each of the three genetic groups.

To adjust the RRMLS, PHol and PHet were transformed into covariables, $\{\varphi_i(\text{PHol})\}$ and $\{\varphi_i(\text{PHet})\}$, respectively, where φ_i is the i^{th} covariable for PHol or PHet. Assuming n knots in point T_i ($i = 1, \dots, n$) and that $T_i \leq t < T_{i+1}$, the covariables may be obtained in the knots i and $i + 1$ through the equations:

$$\varphi_i(t) = \frac{T_{i+1} - t}{T_{i+1} - T_i} \text{ and } \varphi_{i+1}(t) = 1 - \frac{T_{i+1} - t}{T_{i+1} - T_i},$$

respectively. For other values of the T_i and T_{i+1} interval, $\varphi_i(t) = 0$. Assuming that the value found for the genetic group corresponds exactly to the adjusted knot,

Table 1. Crossings between sires and dams of different genetic groups and their respective progenies, as well as number of records for age at first calving (AFC), first lactation length (FLL), and first lactation 305-day milk yield (305MY) of a multiple-breed Holstein (H) x Gyr (G) dairy cattle population.

Sire	Crossing ⁽¹⁾		AFC		FLL		305MY	
	Dam	Progeny	NS ⁽²⁾	NP ⁽³⁾	NS	NP	NS	NP
G	H	1/2HG	163	1,357	167	1,276	172	1,349
Subtotal	-	-	163	1,357	167	1,276	172	1,349
5/8HG	3/8HG	1/2HG	25	36	25	35	25	36
5/8HG	5/8HG	5/8HG	169	1,140	167	1,106	171	1,152
5/8HG	7/8HG	3/4HG	81	455	80	428	82	447
Subtotal	-	-	185	1,631	184	1,569	189	1,635
3/4HG	1/4HG	1/2HG	10	12	9	11	9	11
3/4HG	1/2HG	5/8HG	60	279	58	272	60	286
3/4HG	3/4HG	3/4HG	93	719	88	685	90	718
Subtotal	-	-	103	1,010	98	968	101	1,015
H	G	1/2HG	460	2,490	438	2,295	454	2,459
H	1/4HG	5/8HG	437	1,578	440	1,499	450	1,589
H	1/2HG	3/4HG	684	3,748	668	3,551	693	3,757
Subtotal	-	-	1,008	7,811	980	7,345	1,013	7,805
Total	-	-	1,459	11,809	1,429	11,158	1,475	11,804

⁽¹⁾The fraction in front of the letters HG represents the expected percentage of alleles of the Holstein breed. ⁽²⁾NS, number of sires. ⁽³⁾NP, number of progenies.

$$\varphi_i(T_k) = \begin{cases} 1 & \text{if } i = k \\ 0 & \text{if } i \neq k \end{cases}$$

$$y = X\beta + Za + e$$

The mean trajectories were also adjusted by linear polynomial splines. For PHol, the knots were positioned at the same points defined by the adjusted genetic effects, whereas, for PHet, the knots were positioned at the 0.375, 0.5, and 1.0 points. The effect of dam age at calving was not included in the analysis of AFC. However, the effects of contemporary groups and birth month for AFC or calving month for FLL and 305MY were considered. The random regression model may be presented as:

$$y_{ijklmn} = cg_i + m_j + b_{1(he)} \sum_{he=1}^{H_3} \varphi_{he}(PHet_k) + b_{2(ho)} \sum_{ho=1}^{H_3} \varphi_{ho}(PHol_l) + b_3 age_m + b_4 age_m^2 + \sum_{ho=1}^{H_3} \varphi_{ho}(PHol_l) a_{n(ho)} + e_{ijklmn}$$

where $b_{1(he)}$ and $b_{2(ho)}$ are the regression coefficients associated with the heterozygosis effects and the percentage of alleles of the Holstein breed, respectively; $\varphi_{he}(PHet_k)$ and $\varphi_{ho}(PHol_l)$ are the polynomials associated with knots he and ho, respectively; $a_{n(ho)}$ is the random regression coefficient for the additive genetic effects of animal n associated with knot ho; e_{ijklmn} is the error associated with each observation; and the remaining terms are defined as previously described.

Under a matrix, the regression model may be represented by:

where y is the vector of phenotypes observed; X is the incidence matrix of fixed effects; β is the vector of solutions for fixed effects; Z is the incidence matrix of the polynomial coefficients used to model the direct additive genetic effects; a is the vector of solutions for the random regression coefficients of the direct additive genetic effects; and e is the vector of residues.

The following assumptions were made for the effects included in the MTM model:

$$\begin{bmatrix} \beta_1 & \beta_2 & \beta_3 \end{bmatrix}^t \sim \text{constant};$$

$$\begin{bmatrix} a_1 & a_2 & a_3 \end{bmatrix}^t | A, G_0 \sim N([0 \ 0 \ 0]^t, G_0 \otimes A), \text{ and}$$

$$\begin{bmatrix} e_1 & e_2 & e_3 \end{bmatrix}^t | R_0 \sim N([0 \ 0 \ 0]^t, R_0 \otimes I);$$

where A represents the additive genetic relationship matrix among the animals; G_0 is the matrix of additive genetic (co)variances between traits, so that

$$G_0 = \begin{bmatrix} \sigma^2_1 & \sigma^2_{12} & \sigma^2_{13} \\ \sigma^2_{12} & \sigma^2_2 & \sigma^2_{23} \\ \sigma^2_{13} & \sigma^2_{23} & \sigma^2_3 \end{bmatrix};$$

N is the normal distribution; \otimes is the direct product operator between matrices; R_0 is the matrix of residual variances for the traits, so that

Table 2. Descriptive statistics of age at first calving (AFC), first lactation length (FLL), and first lactation 305-day milk yield (305MY) of a multiple-breed Holstein (H) x Gyr (G) dairy cattle population.

Genetic group ⁽¹⁾	Observations	NCG ⁽²⁾	Mean	Standard deviation	Minimum	Maximum
Age at first calving (day)						
1/2HG	3,895	813	1,024.24	187.40	564.00	1,637.00
5/8HG	2,997	781	1,075.05	192.50	590.00	1,619.00
3/4HG	4,917	832	1,041.57	191.17	574.00	1,611.00
First lactation length (day)						
1/2HG	3,617	753	278.04	94.67	61.00	684.00
5/8HG	2,877	690	274.49	100.41	61.00	695.00
3/4HG	4,664	934	301.70	106.65	61.00	700.00
First lactation 305-day milk yield (kg)						
1/2HG	3,855	788	4,243.47	2,268.63	200.65	15,354.70
5/8HG	3,027	711	3,580.09	1,985.46	200.54	13,182.96
3/4HG	4,922	867	4,497.60	2,079.47	200.54	14,088.00

⁽¹⁾The fraction in front of the letters HG represents the expected percentage of alleles of the Holstein breed. ⁽²⁾NCG, number of contemporary groups.

$$R_0 = \begin{bmatrix} \sigma_{11}^2 & 0 & 0 \\ 0 & \sigma_{22}^2 & 0 \\ 0 & 0 & \sigma_{33}^2 \end{bmatrix};$$

and I is an identity matrix of order equal to the number of observations. For the RRMLS, the following assumptions were made a priori: $\beta \sim \text{constant}$, $a|A, G_0^* \sim N(0, G_0^* \otimes A)$, and $e|R_0^* \sim N(0, R_0^*)$. For both models, inverse-Wishart (IW) distributions were assumed for the matrices of genetic and residual (co)variances, with G_0 and $G_0^* \sim \text{IW}(\Sigma_a^2, n_a)$ and $R_0^* \sim \text{IW}(\Sigma_e^2, n_e)$, where $\Sigma_a^2, \Sigma_e^2, n_a$, and n_e represent the hyperparameters of inverse-Wishart distributions. Information on the complete conditional distributions a priori is available in Sorensen & Gianola (2002).

The samples of the complete conditional distributions of the covariance components were obtained by Gibbs sampler using the GIBBS3F90 software (Misztal et al., 2014). Chains of 1,100,000 iterations were considered, with a burn-in of 100,000 iterations and sampling of (co)variance components for each 100 cycles. Chain size was defined in preliminary analyses, according to the method of Raftery & Lewis (1992), available in the Bayesian output analysis program package (Smith, 2005) of the R software (R Core Team, 2015). The convergence of the chains was evaluated by the criterion proposed by Geweke (1992), available in the same software, and by the visual inspection of the values sampled in each iteration.

Results and Discussion

The RRMLS was considered equivalent to the MTM, since the means of the additive and residual variances obtained by both models were, overall, similar in the three genetic groups evaluated due to the overlap of the high-density intervals (Table 3). This result was already expected, as shown by Misztal (2006).

The values of the deviance information criterion for the MTM and RRMLS were equal to 79.40 and 41.88 for AFC, 43.20 and 29.88 for FLL, and 246.24 and 98.25 for 305MY, respectively, indicating the best fittings for RRMLS. The superiority of the RRMLS in relation to the MTM in terms of statistical adjustment and applicability would be even greater if the number of genetic groups were higher. In the case of the MTM, there could be problems related to the estimability

and convergence of the (co)variance components of the genetic groups that showed few records. The data of these less representative genetic groups could be excluded in order to guarantee estimability and convergence in the MTM. However, in the RRMLS, there is no need to include a knot for each genetic group – the knots may be included only for the genetic group that is considered more representative and of greater interest, or according to statistical criteria. This is an indicative that, if the data were analyzed by the RRMLS, it would not be necessary to exclude the records of the less important genetic groups.

Other works on multiple trait evaluations, carried out with data on beef (Cardoso & Tempelman, 2004; Toral et al., 2012; Lukaszewicz et al., 2015) and dairy (Heins et al., 2011; VanRaden et al., 2011) cattle, showed that variances and heritabilities are specific to each genetic group. Similar results were obtained in the present study, since the additive genetic variances and heritabilities of AFC and 305MY were higher in the 1/2HG groups than in the 5/8HG and 3/4HG groups (Table 3). The lowest averages a posteriori were observed in groups 5/8HG and 3/4HG. These differences may be a consequence of the different processes and intensities of selection to which the different genetic groups were subjected to (Toral et al., 2012).

The intense selection performed on the Holstein breed to increase milk yield (Sørensen et al., 2008; Heins et al., 2011) and the use of a small number of bulls with greater breeding value may have altered the allele frequency in a short period of time (Kim et al., 2013). These authors also verified that there was an increase in the means of the endogamy coefficients of the Holstein breed and a consequent reduction in genetic variability. Therefore, lower values were found for genetic variances, mainly in the groups with a higher fraction of alleles of the Holstein breed (Table 3). It should be noted that the significant differences observed for the variances and heritabilities among groups may be associated with the selective recording of records. In practice, the records of some genetic groups, and not of others, are recorded in the breeding programs of multiple-breed populations. Selective milk recording may also be carried out depending on the yield level of the cow.

The mean a posteriori of the residual variances for AFC, FLL, and 305MY did not differ in the three genetic

groups for both models (Table 3), which was attributed to the overlap of the high-density intervals. This result was a consequence of the formation of contemporary groups, since each group included animals of different racial compositions. In this case, it would be impossible to observe differences in the residual variances due to the environmental variations to which the genetic groups were subjected to. Considering the similarities found for residual variances, the heritabilities of AFC, FLL, and 305MY followed the same trends already described for the additive genetic variances.

Table 3. Posteriori means and limits of the high-density intervals, for 90% of the samples (between parentheses), of the additive (σ_a^2) and residual (σ_e^2) genetic variances and heritabilities (h^2) estimated by the multiple-trait model (MTM) and the random regression model using linear splines (RRMLS) for age at first calving (AFC), first lactation length (FLL), and first lactation 305-day milk yield (305MY) of 1/2 Holstein + 1/2 Gyr (1/2HG), 5/8 Holstein + 3/8 Gyr (5/8HG), and 3/4 Holstein + 1/4 Gyr (3/4HG) cows.

Genetic group	Additive genetic variance	Residual genetic variance	Heritability (h^2)
Multiple-trait model			
AFC ⁽¹⁾			
1/2HG	558 (335; 786)	1,461 (1,268; 1,659)	0.27 (0.17; 0.38)
5/8HG	245 (97; 372)	1,872 (1,716; 2,018)	0.11 (0.05; 0.17)
3/4HG	193 (74; 290)	1,741 (1,625; 1,856)	0.09 (0.04; 0.15)
FLL ⁽²⁾			
1/2HG	1,043 (413; 1,569)	5,605 (5,108; 6,202)	0.15 (0.06; 0.23)
5/8HG	1,051 (607; 1,511)	6,431 (5,903; 6,929)	0.14 (0.08; 0.20)
3/4HG	804 (398; 1,175)	7,880 (7,425; 8,330)	0.09 (0.04; 0.13)
305MY ⁽³⁾			
1/2HG	684 (430; 948)	2,018 (1,795; 2,260)	0.25 (0.16; 0.34)
5/8HG	320 (200; 438)	1,626 (1,492; 1,756)	0.16 (0.10; 0.22)
3/4HG	284 (192; 377)	1,982 (1,871; 2,091)	0.12 (0.08; 0.16)
Random regression model using linear splines			
AFC ⁽¹⁾			
1/2HG	730 (488; 970)	1,448 (1,252; 1,660)	0.33 (0.16; 0.38)
5/8HG	283 (117; 428)	1,862 (1,704; 2,026)	0.13 (0.04; 0.19)
3/4HG	422 (268; 585)	1,627 (1,483; 1,770)	0.20 (0.13; 0.27)
FLL ⁽²⁾			
1/2HG	1,549 (925; 2,139)	5,385 (4,815; 5,927)	0.22 (0.13; 0.30)
5/8HG	864 (528; 1,202)	6,599 (6,161; 7,053)	0.11 (0.07; 0.16)
3/4HG	826 (338; 1,275)	7,893 (7,398; 8,388)	0.09 (0.01; 0.20)
305MY ⁽³⁾			
1/2HG	1,042 (743; 1,324)	1,884 (1,161; 2,080)	0.35 (0.26; 0.44)
5/8HG	356 (227; 488)	1,541 (1,493; 1,764)	0.17 (0.11; 0.24)
3/4HG	362 (237; 491)	1,987 (1,856; 2,112)	0.15 (0.07; 0.16)

⁽¹⁾Variances expressed as days²x10⁻¹. ⁽²⁾Variances expressed as days². ⁽³⁾Variances expressed as kg²x10⁻³.

The correlations between the random regression coefficients used to adjust the breeding values for AFC between groups 1/2HG and 3/4HG and between 5/8HG and 3/4HG did not differ from zero, for both models (Table 4). The same results were obtained for FLL between groups 1/2HG and 3/4HG. These same genetic correlations did not differ from zero when the MTM was adopted.

In the case of 305MY, the correlations between all genetic groups differed from zero when the RRMLS was used, and the highest correlation was found between groups 5/8HG and 3/4HG (Table 4). This means a posteriori of the genetic correlations between groups 1/2HG and 5/8HG and between 1/2HG and 3/4HG were similar. When estimated by the MTM, the correlation between groups 1/2HG and 5/8HG did not differ from zero (Table 4).

The high-density intervals of the correlations that included zero should be highlighted, since the classification of the sires used in the formation of these genotypes may vary according to the genetic group of the targeted female progenies. In this case, distortions may occur when animals with a superior breeding value are chosen as sires and evaluated using data on progenies from different groups. Therefore, predicting breeding values for bulls, according to the genetic composition of progenies, may contribute to improve the quality of the genetic evaluation of multiple-breed populations.

Table 4. A posteriori means and limits of the high-density intervals, for 90% of the samples (between parentheses), of the correlations between the random regression coefficients used to adjust the breeding values for age at first calving (AFC, in days), first lactation length (FLL, in days), and first lactation 305-day milk yield (305MY, in kg) of a multiple-breed Holstein (H) x Gyr (G) dairy cattle population, estimated by the multiple-trait model (MTM) and the random regression model using linear splines (RRMLS).

Genetic group	AFC			FLL			305MY		
	MTM								
1/2HG-5/8HG	0.55	(0.13; 0.96)	0.63	(0.37; 0.87)	0.33	(-0.10; 0.74)			
1/2HG-3/4HG	-0.35	(-0.84; 0.08)	0.10	(-0.83; 0.54)	0.64	(0.40; 0.84)			
5/8HG-3/4HG	-0.15	(-0.84; 0.55)	0.81	(0.67; 0.95)	0.88	(0.80; 0.98)			
RRMLS									
1/2HG-5/8HG	0.55	(0.18; 0.94)	0.77	(0.33; 0.98)	0.52	(0.18; 0.89)			
1/2HG-3/4HG	-0.26	(-0.58; 0.07)	0.09	(-0.29; 0.47)	0.41	(0.14; 0.70)			
5/8HG-3/4HG	0.23	(-0.24; 0.89)	0.52	(0.20; 0.85)	0.87	(0.73; 0.99)			

The breeding value of the three sires with the greater number of female progenies were plotted according with the genetic group of the progenies by which the evaluations were performed (Figure 1), and there were reclassifications. This result confirms those obtained by Martinez et al. (2000), who found that different Gyr

bulls were selected based either on the milk yield of their pure or of their crossbred female progenies.

The differences in the specific combination capacities (Falconer & Mackay, 1996) of each bull may be responsible for part of the interactions observed between the genetic groups of bulls and cows. It should

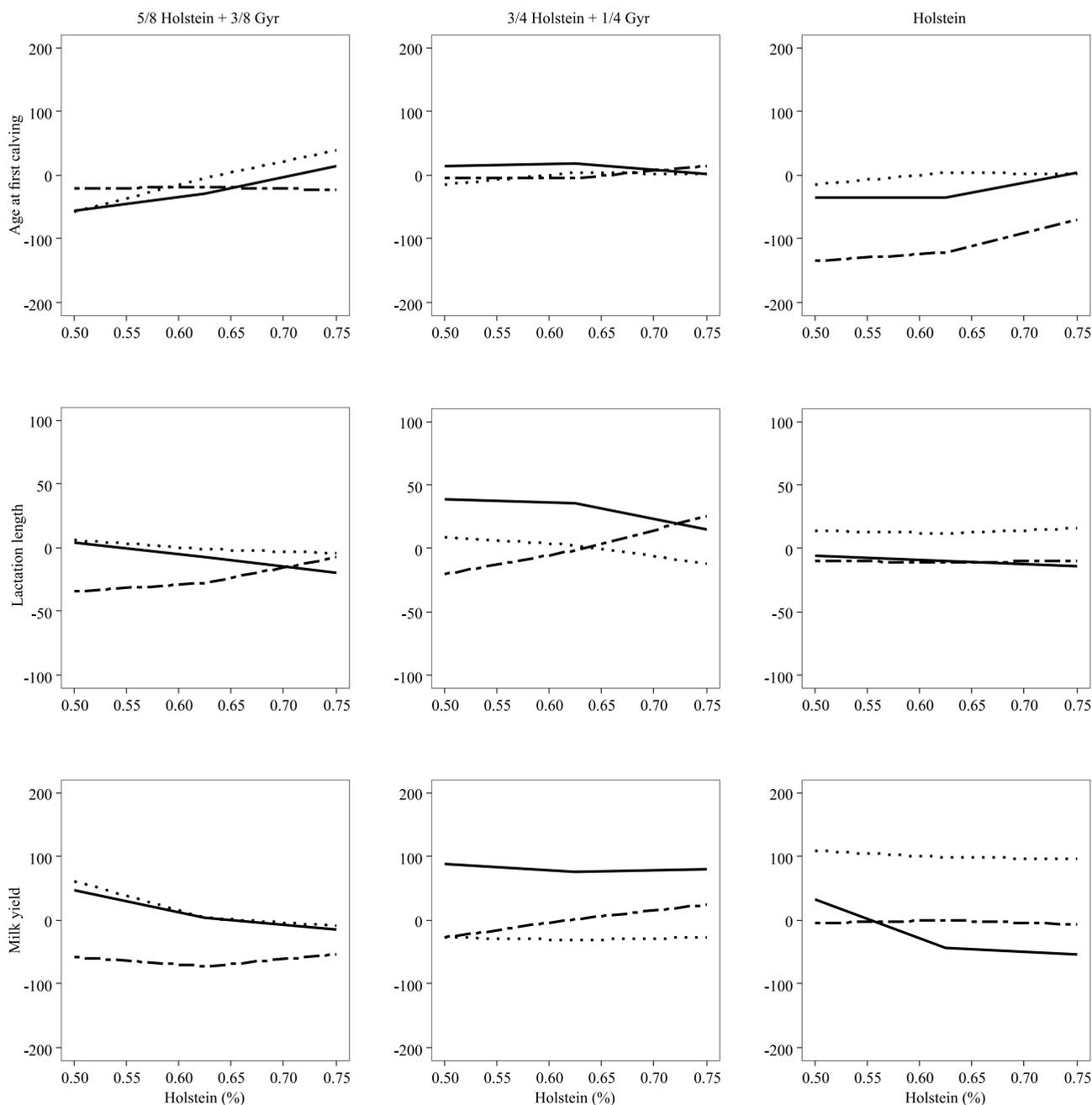


Figure 1. Breeding values for age at first calving (in days), first lactation length (in days), and first lactation 305-day milk yield (in kgx10⁻¹) of three 5/8 Holstein + 3/8 Gyr (5/8HG), 3/4 Holstein + 1/4 Gyr (3/4HG), and pure Holstein sires, which showed the highest number of evaluated female progenies, according to the genetic group of the progenies. Lines represent different sires.

be pointed out that, even when there is an interaction between the breeding value of the bull and the genetic group of the progeny, including yield records of crossbred animals may increase the accuracy of the genetic evaluations of pure-breed sires (VanRaden et al., 2007), as observed for the Holstein and Gyr breeds.

When random regression using linear splines is adopted in the genetic evaluation of a multiple-breed dairy cattle population, it is possible to predict the breeding value of sires and to obtain distinct solutions for each genetic group of progenies. Furthermore, the use of polynomial splines allows simplifying evaluations when knots are precisely adjusted in the genetic groups to be assessed.

Conclusions

1. The random regression model using linear splines, via spline polynomials, is adequate to estimate the genetic parameters for productive and reproductive traits of multiple-breed dairy cattle populations.

2. The breeding value of the sires is affected by the genetic group of the progeny by which it is evaluated.

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